## McGill University ECN 706 Special topics in econometrics Final exam

No documentation allowed Time allowed: 3 hours

10 points

1. Let  $\ell(Y; \theta)$  be the likelihood function for the sample  $Y = (Y_1, \dots, Y_n)'$ . Show that

$$I(\theta) = E\left[-\frac{\partial^2 \log \ell(Y;\theta)}{\partial \theta \partial \theta'}\right].$$

20 points

2. Consider the following assumptions:

H1: the variables  $Y_1, \ldots, Y_n$  are independent and follow the same distribution with density  $f(y; \theta), \theta \in \Theta \subseteq \mathbb{R}^p$ ;

H2: the interior of  $\Theta$  is non-empty, and  $\theta_0$  belongs to the interior of  $\Theta$ ;

H3: the true unknown value  $\theta_0$  is identifiable;

H4: the log-likelihood

$$L_n(y; \theta) = \sum_{i=1}^n \log [f(y_i; \theta)]$$
 is continuous in  $\theta$ ;

H5:  $\mathsf{E}_{\theta_0}[\log f(Y_i; \theta)]$  is finite;

H6: the log-likelihood is such that  $\frac{1}{n}L_n(y;\theta)$  converges almost surely to  $\mathsf{E}_{\theta_0}[\log(Y_i;\theta)]$  uniformly in  $\theta \in \Theta$ ;

H7: the log-likelihood is twice continuously differentiable in open neighborhood of  $\theta_0$ ;

H8:  $I_1(\theta_0) = \mathsf{E}_{\theta_0} \left[ -\frac{\partial^2 \log f(Y;\theta)}{\partial \theta \partial \theta'} \right]$  is finite and invertible.

If  $\hat{\theta}_n$  is consistent sequence of local maxima, show that the asymptotic distribution of  $\sqrt{n}(\hat{\theta}_n - \theta_0)$  is  $N[0, I_1(\theta_0)^{-1}]$ .

20 points

3. State and prove the Neyman-Pearson theorem.

10 points

- 4. Define the following notions:
  - (a) unbiased test;
  - (b)  $\alpha$ -similar test;
  - (c) test with Neyman  $\alpha$ -structure.

10 points

Demonstrate the following relationship between identifiability and unbiased estimation:

if a function  $g(\theta)$  of a parameter  $\theta$  is not identifiable, then there is no unbiased estimator of  $g(\theta)$ .

30 points

6. Consider the standard simultaneous equations model:

$$y = Y\beta + X_1\gamma + u, \qquad (1)$$

$$Y = X_1 \Pi_1 + X_2 \Pi_2 + V, (2)$$

where y and Y are  $T \times 1$  and  $T \times G$  matrices of endogenous variables,  $X_1$  and  $X_2$  are  $T \times k_1$  and  $T \times k_2$  matrices of exogenous variables,  $\beta$  and  $\gamma$  are  $G \times 1$  and  $k_1 \times 1$  vectors of unknown coefficients,  $\Pi_1$  and  $\Pi_2$  are  $k_1 \times G$  and  $k_2 \times G$  matrices of unknown coefficients,  $u = (u_1, \dots, u_T)'$  is a  $T \times 1$  vector of structural disturbances, and  $V = [V_1, \dots, V_T]'$  is a  $T \times G$  matrix of reduced-form disturbances,

$$X = [X_1, X_2]$$
 is a full-column rank  $T \times k$  matrix (3)

where  $k = k_1 + k_2$ . and

$$u$$
 and  $X$  are independent; (4)

$$u \sim N[0, \sigma_u^2 I_T]. \tag{5}$$

- (a) When is the parameter  $\beta$  identified? Explain your answer.
- (b) When is the parameter  $\beta$  weakly identified? Explain your answer.
- (c) Suppose we wish to test the hypothesis

$$H_0(\beta_0): \beta = \beta_0. \tag{6}$$

- i. Describe the standard Wald-type test for  $H_0(\beta_0)$  based on two-stage-least-least squares, and describe its properties.
- ii. Describe an identification-robust procedure for testing  $H_0(\beta_0)$ .
- iii. Discuss the properties of the latter procedure if the model for Y is in fact

$$Y = X_1 \Pi_1 + X_2 \Pi_2 + X_3 \Pi_3 + V \tag{7}$$

where  $X_3$  is a  $T \times k_3$  matrix of fixed explanatory variables.